

Parameter estimation: non-linear least squares and non-linear mixed effects modeling

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Structure

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1. Inverse problems recap

Inverse problem:

- Infer causal factors from observations that produced them
- estimate θ, M to maximize accordance with data
- Assumptions! i.e. usually $\eta_i \sim N(0, \sigma_i^2)$
- error types:



1. Inverse problems recap

writing down Likelihood according to error model:

 $y_i = x_{i|M,\theta} + \eta_i$, $\eta_i \sim N(0, \sigma_i^2)$ additive error

$$L_{y} = \prod_{i=1}^{N} \frac{1}{\sqrt{2 \pi \sigma^{2}}} e^{\frac{-(y_{i} - x_{i})^{2}}{2 \sigma^{2}}}$$

$$l_{y} = \log(\sqrt{2\pi\sigma^{2}}) + \frac{1}{2\sigma^{2}} \left(\sum_{i=1}^{N} -(y_{i} - x_{i})^{2} \right)$$

$$l_y \propto \sum_{i=1}^N (y_i - x_i)^2$$

proportional to least squares problem

- $y_i = x_{i|M,\theta} + \epsilon_i, \qquad \epsilon_i \sim N(0, x_i \sigma_i^2)$

 $y_i = x_{i|M,\theta}(1+\eta_i), \quad \eta_i \sim N(0,\sigma_i^2)$ proportional error

 $l_y \propto \sum_{i=1}^{N} \frac{(y_i - x_i)^2}{x_i}$ proportional to weighted least squares

- pandemia of H1N1 in 2009 (Swineflu)
- children have the highest risk of hospitalization
- used Oseltamivir (Tamiflu) for treatment
- little known about Tamiflu in infants \rightarrow duration of drug therapy?
- virus quantification at Robert-Koch institute, determined from qtip sample
- data points for 36 children, 2 to 5 data points per patient
 - → sparse
- 91 datapoints overall

• assume decay of virus load with treatment to be:

$$V_{estimated}(i,t) = x_0(i)e^{-t CL_v(i)}$$

• which $x_0(i)$ initial viral load, $\frac{copies}{ml}$ $CL_V(i)$ virus clearance, $\frac{1}{day}$ will minimize the error?

•
$$\rightarrow$$
 infer $t_{\mathscr{G}}(i) = \log\left(\frac{x_0(i)}{10}\right) \frac{1}{CL_V(i)}$ time when virus load is unrecognizable

• \rightarrow equal to lower limit of quantification, LLQ = 10

• error models:

 $y_{i} = x_{i|\theta,M} e^{\epsilon_{i}} \text{ exponential } y_{i} = x_{i|\theta,M} (1+\epsilon_{i}) \text{ proportional}$ $\log(y_{i}) = \log(x_{i|\theta,M}) + \epsilon_{i} \qquad \log(y_{i}) = \log(x_{i|\theta,M}) + \log(1+\epsilon_{i})$

- \rightarrow both turn into additive error model when taking logarithm
- fitting:

$$\underset{x_{0}(i),CL_{V}(i) _{t}}{argmin} \sum_{t} \frac{\left(V_{estimated}(i,t) - V_{observed}(i,t)\right)^{2}}{V_{observed}(i,t)}$$
 weighted, or

$$\underset{x_0(i), CL_V(i)}{\operatorname{argmin}} \sum_{t} \left(V_{estimated}(i, t) - V_{observed}(i, t) \right)^2 \quad \text{not weighted}$$

which means



- choices in R:
- nlm: Gauss-Newton type algorithm
- optim: Nelder-Mead, quasi-Newton
- convergence issues:
- \rightarrow use V(i,0) as start value for x0
- multistart for Clv from -10 to 10 in 0.5 sized steps
- choose parameter estimates with minimal objective function value

How to model?

- fit individual data for each patient
- averaging, use mean or median of all data points at each time t
- averaging, use mean or median of virus type grouped data
- use NLME (non linear mixed effect modelling)

• example for fitted curves:



enough data points available

• Optimization landscapes:



influence of weight and grouping on Clv



- number of data points at time 0 for
 - all viruses: 36
 - A sensitive: 18
 - A resistant: 7
 - B sensitive: 11



• example for fitted curves:



- some patients don't behave as expected
- least assumptions made in fitting

• Optimization landscapes:



- ugly landscape, big range of CLv legitimately possible because of sparsity
- assume no error if we fit only 2 points

• Distributions over all patients:



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• influence of noise



• not robust to noise, CLv [0.9, 1.7] \rightarrow treatment time influence $_{18}$

- approach for sparse data
- population model is collection of models of individual observations
- response variability reflects errors and intersubject variability
- N patients, unknown parameters $\theta_{\it pop}$, Ω , σ
- $Y_i = f(x_i | \theta_i) + \epsilon_i(\sigma)$ $\theta_i = \theta_{pop} + \eta_i(\Omega)$
- f nonlinear model
- Y_i partial observations
- Assumptions:
- $\epsilon_i(\sigma)$ measurement errors $i.i.d. \sim N(0, \sigma_i^2) \wedge independent of random effects$
- θ_i random effects ~ $N(\theta_{pop}, \Omega) \wedge independent among groups$

- R: nlme(), but was not documented understandably
- Matlab: nlmefit(), convergence issues → had to use nlmefitsa(), expectation maximization stochastic algorithm

















additive error model + log fit:

- additive error model + log fit:
- very robust to noise
- both random effects go to 0 $\,\rightarrow\,$ try model with only one random effect
- random effect on x0 and CLv: BIC = 462.26
- random effect on x0 only: BIC = 460
- random effect on CLv only: BIC = 476.87
- overall parameters:
- exponential 43345 x0 0.8715 CLv
- proportional 88409 x0 1.1723 CLv
- additive log 50784 x0 0.8646 CLv



6. Summary

- sparse data: fitting to individual patient data makes least assumptions → would be best, but not robust to errors
- fitting on pooled data is robust but doesn't tell us much about single patients
- pooled fitting is a good approximation if we knew what covariate groups data best (i.e. age, virus type, ...)
 BUT
- NLME is the best way to deal with sparse data, robust to errors and keeps characteristics of the groups (patients)
- NLME has easy ways of checking whether it's assumptions are met for the input data
- easy to try out different error models

7. Sources

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